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S. N. Chiu  
*Hong Kong Baptist University, snchiu@hkbu.edu.hk*

H. Y. Lee  
*Hong Kong Baptist University*

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## A regularity condition and strong limit theorems for linear birth–growth processes

By S. N. CHIU AND H. Y. LEE OF HONG KONG

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(Revised Version )

**Abstract.** A linear birth-growth process is generated by an inhomogeneous Poisson process on  $\mathbb{R} \times [0, \infty)$ . Seeds are born randomly according to the Poisson process. Once a seed is born, it commences immediately to grow bidirectionally with a constant speed. The positions occupied by growing intervals are regarded as covered. New seeds continue to form on the uncovered part of  $\mathbb{R}$ . This paper shows that the total number of seeds born on a very long interval satisfies the strong invariance principle and some other strong limit theorems. Also, a gap (an unproved regularity condition) in the proof of the central limit theory in [5] is filled in.

### 1. Introduction

Let  $\Psi \equiv \{(y_i, t_i) \in \mathbb{R} \times [0, \infty)\}$  denote a Poisson point process which is homogeneous in space ( $\mathbb{R}$ ) but can be inhomogeneous in time ( $[0, \infty)$ ). That is, the intensity measure of the process is  $l \times \Lambda$ , where  $l$  is the Lebesgue measure on  $\mathbb{R}$  and  $\Lambda$  an arbitrary locally finite nonzero measure on  $[0, \infty)$ . Consider a linear birth-growth process generated by  $\Psi$  as follows. *Seeds* are *born* according to  $\Psi$ ; once a seed is born, it immediately initiates a bidirectional movement commencing at the location where the seed is born. Each moving frontier progresses with a constant speed  $v > 0$ . The interval passed over by a moving frontier is regarded as covered. New seeds continue to form according to  $\Psi$  only on uncovered intervals in  $\mathbb{R}$  and each seed initiates the same bidirectional grow process. Eventually, the whole line will be completely covered. The resultant structure is known as a one-dimensional Johnson-Mehl tessellation [12], [13] and [16]. Such a linear birth-growth process was studied in details by [4], [5], [6], [8], [9], [11],

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[15] and [17]. Applications can be found in various biological context [1], [7], [18] and [19].

The locations and times of seeds born form a space-time point process  $\Phi$  on  $\mathbb{R} \times [0, \infty)$ , which is a dependently thinned version of the Poisson process  $\Psi$ . The locations of seeds born form a stationary point process on  $\mathbb{R}$  with density

$$\mu \equiv \int_0^\infty \exp \left\{ - \int_0^t 2v(t-u)\Lambda(du) \right\} \Lambda(dt)$$

(see [16]). Throughout the paper we assume that  $\mu < \infty$ .

CHIU [4], HOLST *et al.* [11] and QUINE and ROBINSON [17] established the asymptotic normality of the total number of seeds born on a very long interval. CHIU and QUINE [5] considered a  $d$ -dimensional birth-growth process,  $d \geq 1$ , and obtained the asymptotic normality of the number of seeds born in a very large cube. Moreover, for the case  $d = 1$ , they also generalised the central limit theorem for the total number of seeds formed to a functional central limit theorem (weak invariance principle), which means that the number of seeds born, after suitable normalisation and interpolation, behaves asymptotically like a Brownian motion. Their argument for establishing these weak limit theorems requires the regularity condition that  $\sigma^2 > 0$ , where the parameter  $\sigma^2$  will be explained in Section 2. However, they have only shown numerically that  $\sigma^2 > 0$  for  $\Lambda(dt) = \lambda dt$ ,  $0 < \lambda < \infty$  and  $d = 1, 2, 3$  and 4, but have not considered other  $\Lambda$ . We will fill in this gap for the case  $d = 1$  and generalise the weak invariance principle to the strong invariance principle for the total number of seeds born. Some other strong limit theorems are also shown.

## 2. Regularity condition

CHIU and QUINE [5] established the weak invariance principle for the total number of seeds by considering the total number as the sum of a mixing sequence of random variables. We follow their notation. Denote by  $\xi_z$  the number of seeds born in  $[z, z+1)$ ,  $z \in \mathbb{Z}$ . CHIU and QUINE [5] showed that  $\{\xi_z : z \in \mathbb{Z}\}$  is a stationary  $\alpha$ -mixing sequence. Let

$$S_n \equiv \sum_{0 \leq z \leq n} (\xi_z - \mu)$$

for  $n \in \mathbb{N}$ . Many limit theorems for a mixing sequence require the regularity condition that

$$\sigma^2 \equiv \lim_{n \rightarrow \infty} \frac{\text{var}(S_n)}{n} > 0,$$

which is not necessarily true [2].

CHIU and QUINE [5] derived an expression for  $\sigma^2$  of a  $d$ -dimensional birth-growth process:

$$\sigma^2 = \mu - \int_0^\infty \exp\{-\Delta(t_1)\} \int_0^\infty \omega_d v^d(t_1 + t_2)^d \exp\{-\Delta(t_2)\} \Lambda(dt_2) \Lambda(dt_1)$$

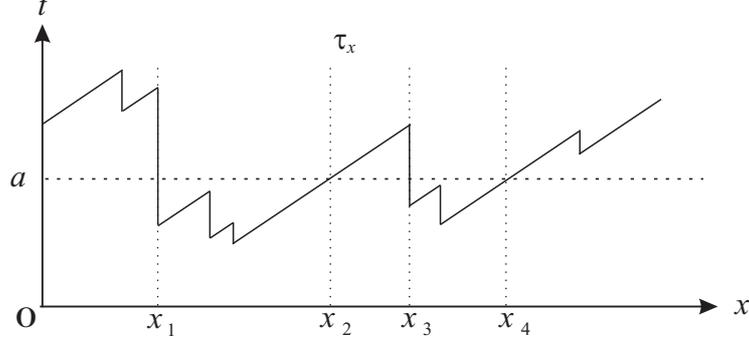


Figure 1: A realisation of the unidirectional process  $\tau_x$ . The path of  $\tau_x$  in  $[x_1, x_2)$ ,  $[x_2, x_3)$  and  $[x_3, x_4)$  are, respectively, in a downblock, upblock and downblock. The upblock and the downblock in  $[x_2, x_4)$  form a complete block. However, the path in  $[0, x_1)$  is neither in an upblock nor in a downblock because  $\tau_0 \neq a$ . In this realisation, the number  $Y_1$  of drops in  $[0, x_2)$  is 4 and the number  $B_2$  of drops in  $[x_2, x_4)$  is 2.

$$\begin{aligned}
& + \int_0^\infty \exp\{-\Delta(t_1)\} \int_0^{t_1} \exp\{\Delta(y)\} \\
& \times \int_y^\infty 2d\omega_d v^d (t_1 + t_2 - 2y)^{d-1} \exp\{-\Delta(t_2)\} \Lambda(dt_2) dy \Lambda(dt_1),
\end{aligned}$$

where  $\Delta(t) = \int_0^{t \vee 0} \omega_d v^d (t - u)^d \Lambda(du)$ ,  $\omega_d = \sqrt{\pi^d} / \Gamma(1 + d/2)$  being the volume of a  $d$ -dimensional sphere of unit radius, and  $x \vee y = \max(x, y)$ . They showed numerically that  $\sigma^2 > 0$  for  $\Lambda(dt) = \lambda dt$ ,  $0 < \lambda < \infty$ , in the cases  $d = 1, 2, 3$  and 4. However, when they discussed the following two models,

$$(2.1) \quad \Lambda(dt) = K t^{j-1} dt, \quad K > 0, j \geq 1,$$

$$(2.2) \quad \Lambda(dt) = \frac{\lambda t^{\alpha-1} e^{-t}}{\Gamma(\alpha)} dt, \quad \lambda > 0, \alpha > 1,$$

they did not show  $\sigma^2 > 0$ . A brute force proof for  $\sigma^2 > 0$ , even if possible, would be extremely laborious. In this paper we present a simple proof for a general  $\Lambda$  in the case  $d = 1$ .

We use the approach of [11] and [18]. Assuming, without loss of generality, that  $v = 1/2$ , we take a shear transformation  $(x, t) \rightarrow (x + t/2, t)$  for the spatio-temporal path of the moving frontiers of the bidirectional growth process. This transformed process can be regarded as the path of the moving frontier of a unidirectional growth process that a seed born initiates not a bidirectional but only a unidirectional movement to the right-hand side with unit speed; the movement stops whenever the frontier meets another seed.

Denote by  $\{\tau_x : x \in \mathbb{R}\}$  the transformed spatial ( $x$ -axis) temporal ( $y$ -axis) stochastic process. See Figure 1. It consists of uniform upward motion at unit speed, representing

the unidirectional movement, with occasional drops that occur wherever the moving frontier meets a seed; the post-drop level is the birth-time of the seed. Let  $x_1 < x_2 < x_3 < \dots$  be the locations such that  $\tau_{x_i} = a$  and  $\tau_x \neq a$  for all  $x \neq x_i$ ,  $i = 1, 2, 3, \dots$ , where  $a > 0$  is an arbitrary fixed time level. Thus,  $x_i$ 's are the location points where  $\{\tau_x\}$  crosses the time level  $a$ . Following the terminology used in [11], we say that the path of  $\{\tau_x\}$  has been separated by  $\{x_1, x_2, \dots\}$  into *upblocks* where  $\tau_x > a$  and *downblocks* with  $\tau_x < a$ . That is to say, the path of  $\{\tau_x\}$  in  $[x_i, x_{i+1})$ ,  $i = 2, 3, \dots$ , is in an upblock if  $\tau_x > a$  for  $(x_i, x_{i+1})$  or in a downblock if  $\tau_x < a$  for  $(x_i, x_{i+1})$ . Note that the path in  $[0, x_1)$  is neither in an upblock nor in a downblock because  $\tau_0 \neq a$  almost surely. The upblocks and downblocks form an alternating sequence. If the block on  $[x_2, x_3)$  is an upblock (such as in Figure 1), then define a complete block as an upblock and the succeeding downblock; otherwise define it as a downblock and the succeeding upblock. Denote by  $B_i$  the total numbers of drops in a complete block from  $x_{2i-2}$  to  $x_{2i}$ ,  $i = 2, 3, \dots$ . The process  $\{\tau_x : x \in \mathbb{R}\}$  is a stationary Markov process, and so the behaviour of the process within each block is independent of its behaviour in other blocks. Therefore, the  $B_i$ 's are independent and identically distributed for all  $i \geq 2$ .

Let  $\alpha_i = x_{2i-2} - x_{2i}$ ,  $i = 2, 3, \dots$ , denote the length of a complete block. HOLST *et al.* [11], p. 911, showed that

$$\mathbf{E}(\alpha_i) = \frac{1}{\Lambda([0, a])} \exp \left\{ \int_0^a (a - u) \Lambda(du) \right\}.$$

The path of  $\tau_x$  in  $[0, x_2)$  is not in a complete block because  $\tau_0 \neq a$ . Let  $\beta_1 = x_2$  denote the length of this incomplete block. Similarly, the path of  $\tau_x$  in  $[x_{2(M_n-1)}, n)$  is not a complete block since  $\tau_n \neq a$ , where  $M_n$  is a random variable such that

$$\{M_n = m\} = \begin{cases} \{\beta_1 + \alpha_2 + \dots + \alpha_{m-1} < n < \beta_1 + \alpha_2 + \dots + \alpha_m\} & m \geq 3, \\ \{\beta_1 < n < \beta_1 + \alpha_2\} & m = 2, \\ \{\beta_1 > n\} & m = 1. \end{cases}$$

That is,  $M_n$  is the total number of complete and incomplete blocks in  $[0, n)$ . Let  $Y_1$  and  $Y_2$  be the numbers of drops in the first and the last incomplete blocks in  $[0, n)$ , respectively. Note that  $Y_2 = 0$  if  $M_n = 1$ . Because of locational stationarity, the distribution is unchanged under the shear transformation (see [11], p. 904 and [18], p. 302) we have

$$(2.3) \quad \text{distribution of } S_n + n\mu = \text{distribution of } Y_1 + \sum_{i=2}^{M_n-1} B_i + Y_2.$$

Denote by  $\mathcal{F}_1 \equiv \sigma\{\beta_1, Y_1\}$  and  $\mathcal{F}_m \equiv \sigma\{\beta_1, \alpha_2, \dots, \alpha_m, Y_1, B_2, \dots, B_m\}$ ,  $m \geq 2$ , the  $\sigma$ -algebras generated by  $\{\beta_1, Y_1\}$  and  $\{\beta_1, \alpha_2, \dots, \alpha_m, Y_1, B_2, \dots, B_m\}$ , respectively. Then  $\{\mathcal{F}_m : m \geq 1\}$  is an increasing sequence of  $\sigma$ -algebras such that  $B_m$  is  $\mathcal{F}_m$ -measurable and independent of  $\mathcal{F}_{m-1}$  for all  $m \geq 2$ , and  $\{M_n = m\} \in \mathcal{F}_m$ .

**Lemma 2.1.** (GUT [10], Theorem 5.3) *Let  $\mathbf{E}(B_2) = \mu_B$  and  $(M_n - 2)_+ = (M_n - 2) \vee 0$ . We have*

$$(2.4) \quad \mathbf{E} \left( \sum_{i=2}^{M_n-1} B_i \right) = \mu_B \mathbf{E}(M_n - 2)_+ = \mu \mathbf{E}(\alpha_2) \mathbf{E}(M_n - 2)_+,$$

$$(2.5) \quad \mathbf{E} \left[ \sum_{i=2}^{M_n-1} B_i - \mu_B(M_n - 2)_+ \right]^2 = \text{var}(B_2) \mathbf{E}(M_n - 2)_+.$$

Since  $\mathbf{E}(Y_1)/n$  and  $\mathbf{E}(Y_2)/n$  vanish as  $n \rightarrow \infty$ , equations (2.3) and (2.4) lead to the result that  $\mathbf{E}(\alpha_2) \mathbf{E}(M_n - 2)_+/n$  tends to 1 as  $n \rightarrow \infty$ .

The variance of  $S_n$  can be expressed as the sum of the variances of  $Y_1$ ,  $\sum_{i=2}^{M_n-1} B_i$  and  $Y_2$  as these random variables are independent. By (2.5), we have

$$\text{var}(S_n) \geq \text{var}(B_2) \mathbf{E}(M_n - 2)_+ + [\mathbf{E}(B_2)]^2 \text{var}(M_n - 2)_+ \geq \text{var}(B_2) \mathbf{E}(M_n - 2)_+.$$

Since  $\{\tau_x : x \in \mathbb{R}\}$  is Markov,  $\text{var}(B_2) > 0$ , and so we obtain:

**Theorem 2.2.**

$$\sigma^2 \equiv \lim_{n \rightarrow \infty} \frac{\text{var}(S_n)}{n} \geq \lim_{n \rightarrow \infty} \frac{\text{var}(B_2) \mathbf{E}(M_n - 2)_+}{n} = \frac{\text{var}(B_2)}{\mathbf{E}(\alpha_2)} > 0.$$

### 3. Strong limit theorems

CHIU and QUINE [5] showed that for  $d = 1$ , if  $\Lambda$  is given by either (2.1) or (2.2) and  $\sigma^2 > 0$ , the  $\alpha$ -mixing coefficient for the sequence  $\{\xi_z : z \geq 0\}$  is

$$\begin{aligned} \alpha(k) &\equiv \sup_{n \in \mathbf{Z}} \{ |\Pr(A_1 \cap A_2) - \Pr(A_1) \Pr(A_2)| : A_1 \in \sigma(\xi_z : z \leq n), \\ &\quad A_2 \in \sigma(\xi_z : z \geq n + k) \} \\ &= O(e^{-k\rho}) \quad \text{as } k \rightarrow \infty, \end{aligned}$$

for some  $\rho > 0$ , and  $S_n$  satisfies the *weak* invariance principle. Using recently published results, we show in the following that for  $d = 1$ ,  $S_n$  satisfies the *strong* invariance principle as well as other strong limit theorems.

**Lemma 3.1.[Complete convergence]** (LIN and LU [14], Theorem 8.5.0) *Let  $1/2 < s \leq 1$ ,  $2 < r \leq \infty$ , and  $1/s < p < r$ . If  $\mathbf{E}(|\xi_1 - \mu|^r) < \infty$  and  $\sum_{k=1}^{\infty} \alpha^{1/\theta}(k) < \infty$  for some  $\theta > [2 + r/(r - p)]ps/(ps - 1)$ , then for every  $\varepsilon > 0$ ,*

$$(3.1) \quad \sum_{k=1}^{\infty} k^{s-2} \Pr \left\{ \max_{1 \leq i \leq k} |S_i| \geq \varepsilon k^s \right\} < \infty.$$

**Lemma 3.2.[Strong law of large numbers]** (CHEN and WU [3]; LIN and LU [14], Remark 8.2.3) *Suppose that  $\varepsilon > 0$ ,  $\sup_z \mathbf{E}(|\xi_z - \mu|^p) < \infty$  for some  $p > 1$ , and*

$$\alpha(k) = \begin{cases} O(k^{-p/(2p-2)-\varepsilon}) & \text{if } 1 < p < 2, \\ O(k^{-2/p-\varepsilon}) & \text{if } p \geq 2. \end{cases}$$

Then

$$(3.2) \quad \frac{S_n - \mathbf{E}(S_n)}{n} = o(1) \quad \text{almost surely.}$$

**Lemma 3.3.**[Strong invariance principle] (LIN and LU [14], Corollary 9.3.1) *If  $\sigma^2 > 0$ ,  $\mathbf{E}|\xi_z - \mu|^r < \infty$  for some  $r > 4$ , and  $\sum_{k=1}^{\infty} \alpha(k)^{1/4-1/r} < \infty$ , then there exists a standard Wiener process  $\{W(x) : x \geq 0\}$  defined on a suitably extended probability space such that*

$$(3.3) \quad S_n - W(n\sigma^2) = O(n^{1/4}(\log n)^{3/2}) \quad \text{almost surely.}$$

**Lemma 3.4.**[Laws of the iterated logarithm] (LIN and LU [14], Corollary 9.3.2) *If  $\sigma^2 > 0$  and for some  $\delta > 0$ ,  $\sup_z \mathbf{E}(|\xi_z - \mu|^{2+\delta}) < \infty$  and  $\alpha(k) = O(k^{-r})$  for  $r > 1 + 2/\delta$ , then*

$$(3.4) \quad \limsup_{n \rightarrow \infty} \frac{S_n}{\sigma \sqrt{2n \log \log n}} = 1 \quad \text{almost surely,}$$

$$(3.5) \quad \liminf_{n \rightarrow \infty} \frac{\max_{1 \leq m \leq n} |S_m|}{\sqrt{\pi^2 n / 8 \log \log n}} = 1 \quad \text{almost surely.}$$

Consider the linear birth-growth processes. For every positive  $\rho$  and  $\theta$ ,  $\sum_{k=1}^{\infty} e^{-k\rho/\theta} < \infty$ . If  $\Lambda$  is locally finite, then  $\mathbf{E}(|\xi_z - \mu|^r) < \infty$  for every  $z \in \mathbf{Z}$  and  $r > 0$ . Thus, we have

**Theorem 3.5.** *If  $\Lambda$  satisfies either (2.1) or (2.2), then equations (3.1)-(3.5) hold.*

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## References

- [1] BENNETT, M. R., and J. ROBINSON: Probabilistic secretion of quanta from nerve terminals at synaptic sites on muscle cells: non-uniformity, autoinhibition and the binomial hypothesis. *Proc. R. Soc. London* **239**, 329–358, 1990
- [2] BRADLEY, R. C.: On the growth of variances in a central limit theorem for strongly mixing sequences, *Bernoulli* **5**, 67–80, 1999
- [3] CHEN, X. P. and Y. H. WU: Strong law for a mixing sequence, *Acta Math. Appl. Sinica* **5**, 367–371, 1989
- [4] CHIU, S. N.: A central limit theorem for linear Kolmogorov's birth-growth models, *Stochastic Process. Appl.* **66**, 97–106, 1997
- [5] CHIU, S. N., and M. P. QUINE: Central limit theory for the number of seeds in a growth model in  $\mathbb{R}^d$  with inhomogeneous Poisson arrivals, *Ann. Appl. Probab.* **7**, 802–814, 1997
- [6] CHIU, S. N., and C. C. YIN: The time of completion of a linear birth-growth model, *Adv. Appl. Probab.* **32**, 620–627, 2000

- [7] COWAN, R., S. N. CHIU and L. HOLST: A limit theorem for the replication time of a DNA molecule, *J. Appl. Probab.* **32**, 296–303, 1995
- [8] ERHARDSSON, T.: On the number of high excursions of linear growth processes, *Stoch. Proc. Appl.* **65**, 31–53, 1996
- [9] GILBERT, E. N.: Random subdivisions of space into crystals, *Ann. Math. Statist.* **33**, 958–972, 1962
- [10] GUT, A.: *Stopped Random Walks. Limit Theorems and Applications.* Springer–Verlag, New York, 1988
- [11] HOLST, L., M. P. QUINE and J. ROBINSON: A general stochastic model for nucleation and linear growth, *Ann. Appl. Probab.* **6**, 903–921, 1996
- [12] JOHNSON, W. A. and R. F. MEHL: Reaction kinetics in processes of nucleation and growth, *Trans. Amer. Inst. Min. Metal. Petro. Eng.* **135**, 416–458, 1939
- [13] KOLMOGOROV, A. N.: On statistical theory of metal crystallisation, *Izvestia Akademii Nauk SSSR, Ser. Mat.* **3**, 355–360, 1937
- [14] LIN, Z. and C. LU: *Limit Theory for Mixing Dependent Random Variables.* Kluwer, Dordrecht, 1996
- [15] MEIJERING, J. L.: Interface area, edge length and number of vertices in crystal aggregates with random nucleation, *Philips Res. Rep.* **8**, 270–290, 1953
- [16] MØLLER, J.: Random Johnson–Mehl tessellations, *Adv. Appl. Probab.* **24**, 814–844, 1992
- [17] QUINE, M. P. and J. ROBINSON: A linear random growth model, *J. Appl. Probab.* **27**, 499–509, 1990
- [18] VANDERBEI, R. J. and L. A. SHEPP: A probabilistic model for the time to unravel a strand of DNA, *Stochastic Models* **4**, 299–314, 1988
- [19] WOLK, C. P.: Formation of one–dimensional patterns by stochastic processes and by filamentous blue–green algae, *Dev. Biol.* **46**, 370–382, 1975

*Department of Mathematics*  
*Hong Kong Baptist University*  
*Kowloon Tong*  
*Hong Kong*  
*E–mail:*  
*snchiu@hkbu.edu.hk*